ON THE PRUSS EXTENSION OF THE HSU-ROBBINS-ERDŐS THEOREM

Soo Hak Sung

ABSTRACT. The Hsu-Robbins-Erdős theorem states that if $\{X_n, n \geq 1\}$ is a sequence of independent and identically distributed random variables, then $EX_1^2 < \infty$ and $EX_1 = 0$ if and only if $\sum_{n=1}^{\infty} P(|\sum_{k=1}^{n} X_k| \geq n\epsilon) < \infty$ for every $\epsilon > 0$. Under some auxiliary conditions, Spătaru (1994) extended this to the case where the X_n are independent, but their distributions come from a finite set. Pruss (1996) proved Spătaru's result under weaker conditions. The purpose of this paper is to improve Pruss conditions.

1. Introduction

Suppose that $N=\{1,2,\cdots\}$ is partitioned into p subsets N_1,\cdots,N_p , and let Y_1,\cdots,Y_p be p random variables. Let $\{X_n,n\geq 1\}$ be a sequence of independent random variables such that X_n has the same distribution as Y_i whenever $n\in N_i, 1\leq i\leq p$, and let $S_n=X_1+\cdots+X_n$. Recently, Spătaru [5] proved that if the auxiliary Conditions A and B(below) hold, then

(1)
$$\sum_{n=1}^{\infty} P(|S_n| \ge \epsilon n) < \infty, \quad \forall \epsilon > 0,$$

if and only if

(2)
$$\sum_{n=1}^{\infty} \sum_{k=1}^{n} P(|X_k| \ge n) < \infty$$

Received May 15, 1998.

¹⁹⁹¹ Mathematics Subject Classification: 60F15.

Key words and phrases: complete convergence, strong law, sums of independent random variables.

The author wishes to acknowledge the financial support of the Korea Research Foundation made in the program year of 1998 (98-015-D00049).

and

(3)
$$\lim_{n \to \infty} \frac{1}{n} \sum_{k=1}^{n} EX_k I(|X_k| < n) = 0.$$

Condition A. For each $i \in \{1, \dots, p\}$, there exists $a_i \in [0, 1]$ and positive constants $C_1(i), C_2(i)$ such that

$$C_1(i)n^{a_i} \le \alpha_i(n) \le C_2(i)n^{a_i+1/2}$$

for all n large, where $\alpha_i(t)=\sharp\{l\in N_i:l\leq t\}$ for $1\leq i\leq p$ and $t\in R$.

Condition **B**. There is a constant C such that , for any $n \in N$ and $i = 1, \dots, p$,

$$\sum_{k \in [n,\infty) \cap N_i} \frac{1}{k^3} \le C \frac{\alpha_i(n)}{n^3}.$$

For the case of p=1, i.e., $\{X_n, n \geq 1\}$ is a sequence of identically distributed random variables, Hsu-Robbins-Erdős theorem (see, Hsu and Robbins[3] and Erdős [1,2]) states that (1) holds if and only if (2) and (3) hold.

For the general case, Spătaru [5] showed that without the auxiliary conditions (1) always implies both (2) and (3). Pruss [4] and Sung [7] independently proved that the converse is not true. On the other hand, Pruss [4] provided some other weaker auxiliary conditions, for every $i \in \{1, \dots, p\}$

(4)
$$[(\mathbf{V}_i \text{ or } \mathbf{W}_i \text{ or } \mathbf{X}_i) \text{ and } \mathbf{Y}_i] \text{ or } \mathbf{Z}_i,$$

under which (2) and (3) imply (1). The reader should be referred to Pruss' paper to find the definitions of the conditions labeled V, W, X and Z. However, Pruss [4] could not prove that (4) can be replaced by Y_i and left it as an open problem. The Condition Y_i is as follows.

Condition Y_i . There is a constant C_i such that

$$\alpha_i(2n) \leq C_i \alpha_i(n)$$

for n sufficiently large.

Note that Condition **B** implies Condition \mathbf{Y}_i for each $i \in \{1, \dots, p\}$, and Condition \mathbf{Z}_i implies \mathbf{Y}_i (see, Pruss [4]).

The purpose of this paper is to solve the open problem. In light of the above note, the following theorem improves not only that of Spătaru [5] but also that of Pruss [4].

THEOREM 1. Assume that Condition Y_i holds for each $i \in \{1, \dots, p\}$. Then (1) holds if and only if (2) and (3) hold.

The proof of the theorem is presented in Section 3. Section 2 collects some preliminary lemmas.

2. Preliminary lemmas

To prove the main theorem we will need the following several lemmas. The next three lemmas come from Pruss [4].

LEMMA 1. Assume that Condition Y_i holds for each $i \in \{1, \dots, p\}$. Then (2) holds if and only if

(5)
$$\sum_{n=1}^{\infty} \sum_{k=1}^{n} P(|X_k| \ge n\delta) < \infty$$

for each $\delta > 0$.

LEMMA 2. Suppose that Y_i has finite mean for each $i \in \{1, \dots, p\}$. Assume that

(6)
$$\lim_{n\to\infty} \frac{1}{n} \sum_{k=1}^n EX_k = 0.$$

Then the following statements are equivalent:

- (i) condition (1) holds, i.e., S_n/n converges completely to 0 as $n \to \infty$.
- (ii) $\sum_{n=1}^{\infty} P(|(S_n^{(i)})^s| \geq n\epsilon) < \infty$, for every $\epsilon > 0$ and each $i \in \{1, \dots, p\}$, where $S_n^{(i)} = \sum_{k \in [1, n] \cap N_i} X_k$ and X^s denotes the symmetrized random variable associated with X.

LEMMA 3. Suppose that Y_i has finite mean for each $i \in \{1, \dots, p\}$. Then (3) and (6) are equivalent.

Now we derive new results under the assumption that Condition \mathbf{Y}_i holds.

LEMMA 4. Assume that Condition Y_i holds. Then there exists a constant c such that

(7)
$$cn\alpha_i(n) \leq \sum_{k=1}^n \alpha_i(k) \leq n\alpha_i(n)$$

for n sufficiently large.

Proof. The second inequality of (7) is obvious, since $\alpha_i(\cdot)$ is a non-decreasing function. On the other hand, we get from the monotonicity of $\alpha_i(\cdot)$ and Condition \mathbf{Y}_i that

$$\begin{split} &\sum_{k=1}^{n} \alpha_{i}(k) \geq \left\lfloor \frac{n}{2} \right\rfloor \alpha_{i} \left(\left\lfloor \frac{n}{2} \right\rfloor \right) \geq \left\lfloor \frac{n}{2} \right\rfloor \frac{1}{C_{i}} \alpha_{i} \left(2 \left\lfloor \frac{n}{2} \right\rfloor \right) \\ &\geq \left\lfloor \frac{n}{2} \right\rfloor \frac{1}{C_{i}} \alpha_{i}(n-1) \geq \left\lfloor \frac{n}{2} \right\rfloor \frac{1}{C_{i}} (\alpha_{i}(n)-1) \geq \left\lfloor \frac{n}{2} \right\rfloor \frac{1}{2C_{i}} \alpha_{i}(n) \end{split}$$

for n sufficiently large, where $\lfloor a \rfloor$ denotes the greatest integer not exceeding a. Thus the first inequality of (7) is proved.

By applying Lemma 4, we have the following lemma which plays an essential role in our main theorem.

LEMMA 5. Assume that Condition \mathbf{Y}_i holds. Then $E|X|\alpha_i(X|) < \infty$ if and only if $\sum_{n=1}^{\infty} \alpha_i(n) P(|X| \ge n) < \infty$.

Proof. Noting that

$$\sum_{n=1}^{\infty} lpha_i(n) P(|X| \geq n) = \sum_{j=1}^{\infty} P(j \leq |X| < j+1) \sum_{n=1}^{j} lpha_i(n),$$

we have by Lemma 4 that

(8) $c_1 \sum_{j=1}^{\infty} j\alpha_i(j) P(j \le |X| < j+1) \le \sum_{n=1}^{\infty} \alpha_i(n) P(|X| \ge n)$ $\le \sum_{i=1}^{\infty} j\alpha_i(j) P(j \le |X| < j+1)$

for some constant c_1 . Also, from the following fact

$$E|X|\alpha_i(|X|) = \sum_{j=0}^{\infty} E|X|\alpha_i(|X|)I(j \le |X| < j+1),$$

we get by Condition \mathbf{Y}_i that

$$\sum_{j=1}^{\infty} j\alpha_{i}(j)P(j \leq |X| < j+1) \leq E|X|\alpha_{i}(|X|)$$

$$(9)$$

$$\leq \sum_{j=0}^{\infty} (j+1)\alpha_{i}(j+1)P(j \leq |X| < j+1)$$

$$\leq c_{2} \sum_{j=1}^{\infty} j\alpha_{i}(j)P(j \leq |X| < j+1)$$

for some constant c_2 . Thus the conclusion follows from (8) and (9). \square

Finally, we will need the following elementary lemma.

LEMMA 6. Assume that Condition Y_i holds. Then

$$\alpha_i(2^j t) \le (2C_i)^j \alpha_i(t)$$

for sufficiently large $t \in R$ and each $j \in N$.

Proof. From the monotonicity of $\alpha_i(\cdot)$ and Condition \mathbf{Y}_i , we have

$$\alpha_i(2t) \le \alpha_i(2(\lfloor t \rfloor + 1)) \le C_i\alpha_i(\lfloor t \rfloor + 1) \le C_i(\alpha_i(t) + 1) \le 2C_i\alpha_i(t)$$

for sufficiently large t. The proof is completed by applying this j-times. \Box

3. Proof of Theorem 1

The fact that (1) implies (2) and (3) was already proved by Spătaru [5]. Conversely, assume that (2) and (3) hold. The proof is based on certain ideas of Stout [6]. In view of Lemma 1, Lemma 2, and Lemma 3, we can assume that for each n, X_n is symmetric. Let $\epsilon > 0$ be given. For each n and $k(1 \le k \le n)$, define $X'_{nk} = X_k I(|X_k| < n^{\rho}), X''_{nk} = X_k I(|X_k| \ge n\epsilon/N)$, and $X'''_{nk} = X_k I(n^{\rho} \le |X_k| < n\epsilon/N)$, where $\rho(0 < \rho < 1)$ and N(positive integer) are chosen later in the proof. Let $S'_n{}^{(i)} = \sum_{k \in [1,n] \cap N_i} X''_{nk}, S''_n{}^{(i)} = \sum_{k \in [1,n] \cap N_i} X''_{nk}$, and $S'''^{(i)} = \sum_{k \in [1,n] \cap N_i} X''_{nk}$.

From the inequality $e^x \leq 1 + x + \frac{x^2}{2}e^{|x|}$ for all $x \in R$, noting that $|X'_{nk}| < n^{\rho}$, we have for t > 0

$$E\left\{\exp\left(\frac{t}{n}X'_{nk}\right)\right\} \le E\left\{1 + \frac{t}{n}X'_{nk} + \frac{t^2}{2n^2}|X'_{nk}|^2 \exp\left(\frac{t}{n}|X'_{nk}|\right)\right\}$$

$$= 1 + E\left\{\frac{t^2}{2n^2}|X'_{nk}|^2 \exp\left(\frac{t}{n}|X'_{nk}|\right)\right\}$$

$$\le 1 + \frac{t^2}{2n^2} \exp\left(\frac{t}{n^{1-\rho}}\right) E|X'_{nk}|^2$$

$$\le \exp\left\{\frac{t^2}{2n^2} \exp\left(\frac{t}{n^{1-\rho}}\right) E|X'_{nk}|^2\right\}.$$

By independence

$$\begin{split} E\left\{\exp\left(\frac{t}{n}S_n'^{(i)}\right)\right\} &= \prod_{k\in[1,n]\cap N_i} E\left\{\exp\left(\frac{t}{n}X_{nk}'\right)\right\} \\ &\leq \exp\left\{\frac{t^2}{2n^2}\alpha_i(n)\exp\left(\frac{t}{n^{1-\rho}}\right)EY_i^2I(|Y_i|< n^\rho)\right\} \\ &\leq \exp\left\{\frac{t^2}{2n^{1-\rho}}\exp\left(\frac{t}{n^{1-\rho}}\right)E|Y_i|\right\}. \end{split}$$

Hsu-Robbins-Erdős theorem

Thus, choosing $t = 2 \log n/\epsilon$, we have

$$\begin{split} P(S_n^{\prime}{}^{(i)} &\geq n\epsilon) \leq e^{-t\epsilon} E\left\{\exp\left(\frac{t}{n} S_n^{\prime}{}^{(i)}\right)\right\} \\ &\leq \exp\left\{-t\epsilon + \frac{t^2}{2n^{1-\rho}} \exp\left(\frac{t}{n^{1-\rho}}\right) E|Y_i|\right\} \\ &\leq \frac{2}{n^2} \end{split}$$

for n sufficiently large, since $\frac{t^2}{n^{1-\rho}} \exp(\frac{t}{n^{1-\rho}}) \to 0$ as $n \to \infty$. Thus

(10)
$$\sum_{n=1}^{\infty} P(S_n^{\prime}{}^{(i)} \ge n\epsilon) < \infty.$$

From the definition of X''_{nk} , it follows that

$$P(S_n''^{(i)} \ge n\epsilon) \le \sum_{k \in [1, n] \cap N_i} P\left(|X_k| \ge \frac{n\epsilon}{N}\right).$$

This entails by Lemma 1 that

(11)
$$\sum_{n=1}^{\infty} P(S_n''^{(i)} \ge n\epsilon) < \infty.$$

Since $|X_{nk}'''| < n\epsilon/N, S_n'''^{(i)} \ge n\epsilon$ implies that there are at least N non-zero X_{nk}''' for $k \in [1, n] \cap N_i$. Hence

$$\begin{split} P(S_n^{\prime\prime\prime(i)} \geq n\epsilon) &\leq \sum_{k_1 < \dots < k_N} P(X_{n,k_1}^{\prime\prime\prime} \neq 0, \dots, X_{n,k_N}^{\prime\prime\prime} \neq 0) \\ &\leq \sum_{k_1 < \dots < k_N} P(|X_{k_1}| \geq n^{\rho}, \dots, |X_{k_N}| \geq n^{\rho}) \\ &= \binom{\alpha_i(n)}{N} [P(|Y_i| \geq n^{\rho})]^N \\ &\leq \binom{\alpha_i(n)}{N} [P(|Y_i|\alpha_i(|Y_i|) \geq n^{\rho}\alpha_i(n^{\rho}))]^N, \end{split}$$

where the summation $\sum_{k_1 < \dots < k_N}$ is taken for all N-tuples (k_1, \dots, k_N) with $k_1 < \dots < k_N$ and $k_j \in [1, n] \cap N_i$ for each j. By Markov's inequality,

$$P(S_n'''^{(i)} \ge n\epsilon) \le {\alpha_i(n) \choose N} \left[\frac{E|Y_i|\alpha_i(|Y_i|)}{n^\rho \alpha_i(n^\rho)} \right]^N$$

$$\le [E|Y_i|\alpha_i(|Y_i|)]^N \frac{\alpha_i^N(n)}{n^{\rho N} \alpha_i^N(n^\rho)}.$$

Thus we have

$$\begin{split} \sum_{n=1}^{\infty} P(S_n'''^{(i)} \geq n\epsilon) &\leq [E|Y_i|\alpha_i(|Y_i|)]^N \sum_{n=1}^{\infty} \frac{\alpha_i^N(n)}{n^{\rho N} \alpha_i^N(n^{\rho})} \\ &= [E|Y_i|\alpha_i(|Y_i|)]^N \sum_{k=0}^{\infty} \sum_{n=2^k}^{2^{k+1}-1} \frac{\alpha_i^N(n)}{n^{\rho N} \alpha_i^N(n^{\rho})} \\ &\leq [E|Y_i|\alpha_i(|Y_i|)]^N \sum_{k=0}^{\infty} \frac{\alpha_i^N(2^{k+1})}{2^{k(\rho N-1)} \alpha_i^N(2^{k\rho})} \\ &\leq c[E|Y_i|\alpha_i(|Y_i|)]^N \sum_{k=0}^{\infty} \frac{(2C_i)^{N(k(1-\rho)+2)}}{2^{k(\rho N-1)}}, \end{split}$$

for some constant c > 0. The last inequality follows from Lemma 6, since

$$\begin{split} \alpha_i(2^{k+1}) &\leq \alpha_i(2^{\lfloor k(1-\rho)\rfloor+2}2^{k\rho}) \\ &\leq (2C_i)^{\lfloor k(1-\rho)\rfloor+2}\alpha_i(2^{k\rho}) \\ &\leq (2C_i)^{k(1-\rho)+2}\alpha_i(2^{k\rho}) \end{split}$$

for k sufficiently large. Choose $0 < \rho < 1$ such that $(2C_i)^{1-\rho}/2^{\rho} < 1$ and then choose large integer N such that $2((2C_i)^{1-\rho}/2^{\rho})^N < 1$. Then we have, noting that $E|Y_i|\alpha_i(|Y_i|) < \infty$ by Lemma 5, that

(12)
$$\sum_{n=1}^{\infty} P(S_n^{\prime\prime\prime(i)} \ge n\epsilon) < \infty.$$

From (10), (11), and (12), we get that

(13)
$$\sum_{n=1}^{\infty} P(S_n \ge 3p\epsilon n) < \infty,$$

since

$$\sum_{n=1}^{\infty} P(S_n \ge 3p\epsilon n)$$

$$\le \sum_{n=1}^{\infty} \sum_{i=1}^{p} [P(S_n^{\prime(i)} \ge n\epsilon) + P(S_n^{\prime\prime(i)} \ge n\epsilon) + P(S_n^{\prime\prime\prime(i)} \ge n\epsilon)].$$

By replacing X_k by $-X_k$ in (13), we obtain that

(14)
$$\sum_{n=1}^{\infty} P(S_n \le -3p\epsilon n) < \infty.$$

Combining (13) and (14) gives

$$\sum_{n=1}^{\infty} P(|S_n| \ge 3p\epsilon n) < \infty,$$

and so (1) is proved.

References

- P. Erdős, On a theorem of Hsu and Robbins, Ann. Math. Statist. 20 (1949), 286-291.
- [2] _____, Remark on my paper "On a theorem of Hsu and Robbins", Ann. Math. Statist. 21 (1950), 138.
- [3] P. L. Hsu and H. Robbins, Complete convergence and the law of large numbers, Proc. Nat. Acad. Sci. U.S.A. 33 (1947), 25-31.
- [4] A. R. Pruss, On Spătaru's extension of the Hsu-Robbins-Erdős law of large numbers, J. Math. Anal. Appl. 199 (1996), 558-578.
- [5] A. Spătaru, A generalization of the Hsu-Robbins-Erdős theorem, J. Math. Anal. Appl. 187 (1994), 371-383.

Soo Hak Sung

- [6] W. F. Stout, Some results on the complete and almost sure convergence of linear combinations of independent random variables and martingale differences, Ann. Math. Statist. 39 (1968), 1549-1562.
- [7] S. H. Sung, A note on Spătaru's complete convergence theorem, J. Math. Anal. Appl. 206 (1997), 606-610.

Department of Applied Mathematics, Pai Chai University, Taejon 302-735, Korea

E-mail: sungsh@woonam.paichai.ac.kr