



$$P_{fa} = \prod_{i=1}^K \left(1 + \frac{q}{M+1-i}\right)^{-1} \quad (2.2b)$$

The  $S_c^{P_{fa}}$  can be calculated using (2.2).

$$S_c^{P_{fa}} = -q \ln q \sum_{i=1}^K \frac{1}{M+1-i} \left(1 + \frac{q}{M+1-i}\right)^{-1} \quad (2.3)$$

If M and K are given, the q can be calculated for preassigned value  $P_{fa}$  using (2.2b). And using this value, we can calculate the  $S_c^{P_{fa}}$ . Since q is dependent only on  $P_{fa}$ ,  $S_c^{P_{fa}}$  is also dependent only on  $P_{fa}$ , i.e.  $S_c^{P_{fa}}$  is independent of c, the weibull skewness parameter. Therefore we can plot  $S_c^{P_{fa}}$  versus  $P_{fa}$ . Figure 2 is the plot of q versus  $P_{fa}$ . And figure 3 is the plot of  $S_c^{P_{fa}}$ .

In figure 3 we can see that (1) if c is overestimated,  $P_{fa}$  will increase (2) This increase is large for small value of  $P_{fa}$ . For example  $S_c^{P_{fa}} = -21.51$  for  $P_{fa} = 10^{-5}$ . This means  $P_{fa} \approx c^{-21.51}$  at  $P_{fa} = 10^{-5}$ . If the true value of c is actually 85 % of estimated value,  $P_{fa}$  increases approximately thirty three times. In other words,  $P_{fa}$  becomes  $3.3 \times 10^{-4}$ . This is rather large performance degradation. The threshold multiplier is related to average decision threshold (ADT) [6]. Its formal definition is  $ADT = E(\alpha Z) / \mu$ , where  $\alpha$  is threshold multiplier which is function of estimation method and preassigned  $P_{fa}$ , the random variable Z is the result of the estimation method used in CFAR system,  $\mu$  is the mean clutter power level. E denotes expectation operation. Actually the ADT is the required SNR for  $P_d = 0.5$  for low  $P_{fa}$ . The smaller the threshold multiplier becomes, the smaller becomes the CFAR loss. Therefore if we overestimate c and we are to maintain the preassigned  $P_{fa}$ , the CFAR detector will increase q, results in increase of  $\alpha$ . This will increase CFAR loss. Therefore accurate estimation of skewness parameter is crucial for performance of CFAR detector.

### 3. MAP estimator for the Weibull clutter

#### 3.1 Motivation

The skewness parameter of the Weibull clutter is varying as grazing angle ( $\phi$ ) varies.

The table 1 (figure 1 of [2]) shows the variation of sea clutter skewness parameter for several grazing angle.

grazing angle( $\phi$ degree)	skewness parameter(c)
1	1.16
5	1.65
30	1.78

table 1 skewness parameter variation

Test condition : (1) sea state 3 (2)  $K_u$  band (3) Horizontal polarization (4) 0.1  $\mu$ S pulse width

The skewness parameter can not be known exactly due to conditions such as weather. But we can know the approximate value of the skewness parameter as the grazing angle varies. In [1] the ML (maximum likelihood) estimator for the mean power of Weibull clutter was proposed. The assumption that the scale parameter can be known accurately is not realistic. In [5], R. Ravid et al derived ML estimator for both scale and skewness parameter assuming that no information of the scale parameter is available. This ML estimator is proved to be CFAR, but this ML estimator doesn't use any prior information of scale parameter. In this paper, we assume that partial information of scale parameter is known and propose MAP estimator for the Weibull clutter. Using MAP estimator we can design CFAR detector. It can be shown that the OW estimator proposed in [1] and ML estimator proposed in [5] is special case of proposed estimator.

#### 3.2 Derivation of the MAP estimator for weibull clutter.

Firstly we assume as following.

#### Assumption

1. Linear law detector is used.
2. The weibull samples are i.i.d
3. The pdf of the skewness parameter is given by  $f_c(c)$

The conditional pdf  $f_{X|c}(X|c)$  is given by

$$f_{X|c}(X|c) = \prod_{i=1}^M p_L(x_i, c, b) \quad (3.4)$$

$$= \left(\frac{c}{b^c}\right)^M \prod_{i=1}^M x_i^{c-1} \exp\left(-\frac{x_i^c}{b^c}\right)$$

Let's define likelihood function as

$$L(X, c, b) = \ln(f_c(c) f_{X|c}(X|c)) \quad (3.5)$$

$$= \ln\left(f_c(c) \prod_{i=1}^M p_L(x_i, c, b)\right)$$

$$= \ln\left(f_c(c) \left(\frac{c}{b^c}\right)^M \prod_{i=1}^M x_i^{c-1} \exp\left(-\frac{x_i^c}{b^c}\right)\right)$$

Differentiate (3.5) with respect to b and equating zero, we obtain

$$\frac{\partial L}{\partial b} = 0 : \hat{b} = \left[\frac{1}{M} \sum_{i=1}^M x_i^{\frac{c}{b^c}}\right]^{\frac{1}{c}} \quad (3.6)$$

Differentiate (3.5) with respect to c and equating zero, we obtain

$$\frac{\partial L}{\partial c} = \frac{1}{f_c(c)} \frac{\partial f_c(c)}{\partial c} + \frac{M}{c} + \sum_{i=1}^M \ln x_i - \left[ \sum_{i=1}^M \left(\frac{x_i}{b}\right)^c \ln x_i \right] = 0 \quad (3.7)$$

If we assume the pdf of skewness parameter is gaussian pdf with mean  $\mu_c$  and standard deviation  $\sigma_c$ , eq (3.7) becomes

$$\frac{\partial L}{\partial c} = -\frac{(\hat{c} - \mu_c)}{\sigma_c^2} + \frac{M}{c} + \sum_{i=1}^M \ln x_i - \left[ \sum_{i=1}^M \left(\frac{x_i}{b}\right)^{\hat{c}} \ln x_i \right] = 0 \quad (3.8)$$

Substituting  $\hat{b}$  in (3.8) using (3.6), eq (3.8) becomes,

$$\frac{\partial L}{\partial c} = -\frac{(\hat{c} - \mu_c)}{\sigma_c^2} + \frac{M}{c} + \sum_{i=1}^M \ln x_i - M \left[ \frac{\sum_{i=1}^M x_i^{\hat{c}} \ln x_i}{\sum_{i=1}^M x_i^{\hat{c}}} \right] = 0 \quad (3.9)$$

The MAP estimate,  $\hat{c}$  for the skewness parameter of the Weibull clutter can be obtained solving eq (3.9). Then  $\hat{b}$  can be obtained using eq (3.6). If no prior information is available, we can think of c as uniformly distributed. Then  $\frac{\partial f_c(c)}{\partial c} = 0$ . Therefore eq

(3.9) becomes

$$\frac{M}{c} + \sum_{i=1}^M \ln x_i - M \left[ \frac{\sum_{i=1}^M x_i^{\hat{c}} \ln x_i}{\sum_{i=1}^M x_i^{\hat{c}}} \right] = 0 \quad (3.10)$$

Eq (3.6) and eq (3.9) are used to design two parameter ML CFAR [5].

As  $\sigma_c$  goes to zero,  $\hat{c} \rightarrow \mu_c$ . Then eq (3.6) becomes

$$\hat{b} = \left[ \frac{1}{M} \sum_{i=1}^M x_i^{\mu_c} \right]^{\frac{1}{\mu_c}} \quad (3.11)$$

Eq (3.11) can be used to design parametric one parameter OW CFAR [1].

Remark : If we assume the more plausible a posteriori probability  $f_c(c)$ , the more accurate estimate of the skewness parameter we can obtain.

#### 4. Simulation result

##### 4.1 Performance Comparison between ML and MAP estimator.

Since the closed form solution of  $P_{fa}$  for MAP detector is hard to find, we resort to statistical simulation method. The  $P_{fa}$  of the MAP detector is given by

$$P_{fa} = P(y > T) = P(y > \delta \alpha^{\frac{1}{c}}) \quad (4.1)$$

where y is the sample in test cell and  $\delta$  and  $\hat{c}$  is estimated parameters using M weibull samples. If we simulate an event with

probability  $p$  with  $m$  trials ,the accuracy of simulation is approximately given by

$$\frac{\text{std}(\hat{p})}{\hat{p}} \approx \frac{1}{\sqrt{mp}} \quad (4.2)$$

If 5 % error is to be guaranteed,  $m$  should be

$$m = \frac{400}{p} \quad (4.3)$$

If we simulate  $p = 10^{-5}$ ,  $m = 4 \cdot 10^7$ .The estimated  $P_{fa}$  is obtained by

$$\hat{P}_{fa} = \frac{1}{m} \sum_{j=1}^m P(y > T_j) = \frac{1}{m} \sum_{j=1}^m \exp\left(-\left(\frac{T_j}{b}\right)^c\right) \quad (4.4)$$

where  $T_j = \hat{b}_j \frac{1}{c}$ .

$b$  and  $c$  : The  $b$  and  $c$  used in generating weibull sample.

Algorithm used in obtaining  $\hat{P}_{fa}$ .

for  $j = 1$  to  $m$ , do

step 1. given  $\alpha$ , generate  $M$  weibull samples using predetermined  $b$  and  $c$ .

step 2. calculate  $\hat{b}_j, \hat{c}_j$  therefore  $T_j$  according to specified CFAR method.

step 3.  $P_{fa}(j) = \exp\left(-\left(\frac{T_j}{b}\right)^c\right)$

end for loop.

step 4  $\hat{P}_{fa} = \frac{1}{m} \sum_{j=1}^m P_{fa}(j)$

We calculated the  $P_{fa}$  of ML CFAR ( $\sigma_c = \infty$  case) MAP CFAR ( $\sigma_c = 0.5, \sigma_c = 0.2, \sigma_c = 0.1$  case), Parametric ML CFAR ( $\sigma_c = 0$  case ).In figure 4, the original weibull samples has parameter  $b=3$ ,  $c=1$  and the number of trial ( $m$ ) = 10000. We can see that as  $\sigma_c$  decreases, the threshold multiplier decreases. So does the CFAR loss. And we can conclude that the proposed MAP detector performs better than ML detector for uniform weibull clutter.

## 6. Conclusion

Firstly, we considered the sensitivity analysis of OS CFAR. From this analysis we concluded that the accurate estimation of skewness parameter is crucial for the performance of CFAR detector. We proposed the MAP CFAR algorithm ,which have smaller CFAR loss than ML algorithm. The proposed algorithm is flexible since experimental result (the pdf of skewness parameter) can be embedded into the algorithm. And this algorithm was shown to be the general case of parametric ML detector (the case that full information of skewness is known) and nonparametric ML detector (the case that no information of skewness is known).

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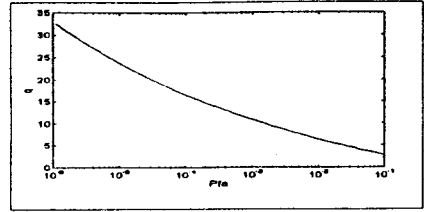


Figure 2.  $q$  value as a function of  $P_{fa}$  (OS CFAR)

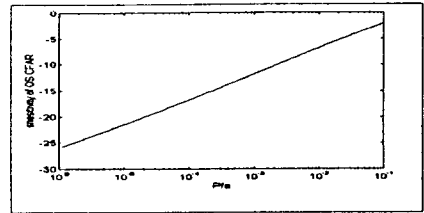


Figure 3. Sensitivity of OS CFAR as a function of  $P_{fa}$

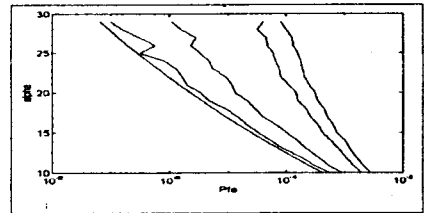


figure 4. Performance comparison between ML CFAR and MAP CFAR ( from top to bottom ML CFAR, MAP CFAR  $\sigma_c = 0.5, \sigma_c = 0.2, \sigma_c = 0.1$  ideal case)