

# GEOMETRIC ERGODICITY AND EXISTENCE OF HIGHER-ORDER MOMENTS FOR DTARCH( $p, q$ ) PROCESS<sup>†</sup>

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## ABSTRACT

We consider a double threshold AR-ARCH type process and give sufficient conditions under which the higher-order moments exist. Geometric ergodicity and strict stationarity are also studied.

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*Keywords.* Double threshold AR-ARCH model, irreducible, geometric ergodicity, stationarity, moment.

## 1. INTRODUCTION

Threshold autoregressive models of Tong (1978) and Tong and Lim (1980) are quite popular in the nonlinear time series literature. ARCH (autoregressive conditional heteroscedasticity) process was proposed by Engle (1982) to explain time series with conditional heteroscedastic variances. Tong (1990) suggested a threshold model with an ARCH error (SETAR-ARCH) which combines the advantage of the threshold AR model and ARCH model. Double threshold AR-ARCH model is a natural extension of SETAR-ARCH model. Family of ARCH model has proven useful in financial applications and have great attention in economics and statistical literature (see, *e.g.*, Engle and Bollerslev, 1986; Bollerslev *et al.*, 1992; Schwert, 1989; Nelson, 1991; Guégan and Diebolt, 1994; Li and Li, 1996; Wong and Li, 1997). Stationarity, geometric ergodicity (or strongly mixing/absolute regularity) and existence of higher-order moments are of importance in the statistical inference of model. There are many papers to discuss estimations of coefficients and threshold for threshold AR-ARCH models and

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distribution theory for the estimators. To study the limiting distribution of the estimator, it is assumed, in general, the absolute regularity and higher order moment conditions (see Chan, 1993; Masry and Tjøstheim, 1995; Li and Li, 1996; Lu, 1996; An *et al.*, 1997; Hansen, 1999 and 2000; Ling, 1999; Hwang and Woo, 2001; Lee, 2002).

In this paper, we consider the double threshold autoregressive model with ARCH errors (DTARCH( $p, q$ ))  $\{y_t\}$  given by

$$y_t = \sum_{i=1}^r \left( \theta_{i0} + \sum_{j=1}^p \theta_{ij} y_{t-j} \right) I_{it} + \varepsilon_t, \quad (1.1)$$

$$\varepsilon_t = \sqrt{h_t} e_t, \quad (1.2)$$

$$h_t = \sum_{k=1}^s \left( \alpha_{k0} + \sum_{j=1}^q \alpha_{kj} \varepsilon_{t-j}^2 \right) J_{kt}, \quad (1.3)$$

where  $\{e_t\}$  is a sequence of independent and identically distributed (*iid*) random variables with mean 0 and variance  $\sigma^2$ ,  $I_{it} = I(a_{i-1} \leq y_{t-b} < a_i)$  and  $J_{kt} = I(b_{k-1} \leq y_{t-d} < b_k)$ ,  $-\infty = a_0 < a_1 < \dots < a_r = \infty$ ,  $-\infty = b_0 < b_1 < \dots < b_s = \infty$ ,  $b, d \in \{1, 2, \dots, p+q\}$ ,  $\alpha_{k0} > 0$ ,  $\alpha_{kj} \geq 0$ ,  $i = 1, 2, \dots, r$ ,  $j = 1, 2, \dots, q$ ,  $k = 1, 2, \dots, s$ , and  $I(A)$  denotes the indicator function of a set  $A$ .

Goal of this paper is to find some region of coefficients on which  $y_t$  is strictly stationary and geometrically ergodic. Existence of higher-order moments is also considered.

In Section 2, we consider the geometric ergodicity and 4<sup>th</sup> moment problem. Strict stationarity and higher-order moment condition are studied in Section 3.

## 2. GEOMETRIC ERGODICITY AND MOMENTS

In this section, we define a Markov chain  $Y_t$  by

$$Y_t = (y_t, y_{t-1}, \dots, y_{t-p-q+1}). \quad (2.1)$$

$y_t$  given by (1.1)–(1.3) can be rewritten as

$$y_t = g_1(y_{t-1}, \dots, y_{t-p}) + g_2(y_{t-1}, \dots, y_{t-p-q}) e_t. \quad (2.2)$$

Here

$$g_1(y_{t-1}, \dots, y_{t-p}) = \sum_{i=1}^r \left( \theta_{i0} + \sum_{j=1}^p \theta_{ij} y_{t-j} \right) I_{it}, \quad (2.3)$$

$$g_2^2(y_{t-1}, \dots, y_{t-p-q}) = \sum_{k=1}^s \left( \alpha_{k0} + \sum_{j=1}^q \alpha_{kj} \varepsilon_{t-j}^2 \right) J_{kt}, \tag{2.4}$$

where

$$\varepsilon_{t-j} = y_{t-j} - \sum_{i=1}^r \left( \theta_{i0} + \sum_{l=1}^p \theta_{il} y_{t-j-l} \right) I_{i(t-j)}, \quad j = 1, \dots, q.$$

Throughout this paper, we assume that

$$\theta_j = \max_{1 \leq i \leq r} |\theta_{ij}|, \quad 0 \leq j \leq p, \tag{2.5}$$

$$\alpha_l = \max_{1 \leq k \leq s} \alpha_{kl}, \quad 0 \leq l \leq q. \tag{2.6}$$

LEMMA 2.1. Let  $v(z) = \sum_{i=1}^n \gamma_i z_i^m$  and  $z = (z_1, \dots, z_n)$  for  $n, m \in \mathbb{Z}^+$ . If  $\sum_{i=1}^n \xi_i < 1$ , we may choose  $\gamma_i > 0$ ,  $i = 1, \dots, n$  so that for some positive constant  $r < 1$ ,

$$\gamma_1 \left( \sum_{i=1}^n \xi_i z_i^m \right) + \sum_{i=2}^n \gamma_i z_{i-1}^m \leq r v(z). \tag{2.7}$$

PROOF. Let  $\delta > 0$  be such that  $\sum_{i=1}^n \xi_i + \delta = 1$ . Choose  $\gamma_1 > 0$  arbitrary and define

$$\gamma_{i+1} = \gamma_1 \left( 1 - \xi_1 - \dots - \xi_i - \frac{i\delta}{n} \right), \quad i = 1, 2, \dots, n-1. \tag{2.8}$$

Then following two inequalities hold:

$$\gamma_1 \xi_i + \gamma_{i+1} \leq \gamma_i \left( 1 - \frac{\delta}{n} \right), \quad 1 \leq i \leq n-1, \tag{2.9}$$

$$\gamma_1 \xi_n \leq \gamma_n \left( 1 - \frac{\delta}{n} \right). \tag{2.10}$$

Using (2.9) and (2.10), we have that

$$\sum_{i=1}^{n-1} (\gamma_1 \xi_i + \gamma_{i+1}) z_i^m + \gamma_1 \xi_n z_n^m \leq \left( 1 - \frac{\delta}{n} \right) \sum_{i=1}^n \gamma_i z_i^m. \tag{2.11}$$

Thus (2.7) holds with  $r = (1 - \delta/n)$ . □

ASSUMPTION 1.1. The *iid* random variables  $\{e_t\}$  have a probability density function that is positive over  $R^1$ .

ASSUMPTION 1.2.  $\theta_j$  and  $\alpha_l$  ( $1 \leq j \leq p$ ,  $1 \leq l \leq q$ ) hold the following:

- (a)  $(\sum_{j=1}^p \theta_j)^2 + (1 + \sum_{j=1}^p \theta_j)^2 (\sum_{l=1}^q \alpha_l) < 1$ ;  
 (b)  $(1 + 3\sigma^2)(\sum_{j=1}^p \theta_j)^4 + (Ee_t^4 + 3\sigma^2)(1 + \sum_{j=1}^p \theta_j)^4 (\sum_{l=1}^q \alpha_l)^2 < 1$ .

**THEOREM 2.1.** *Let Assumption 1.1 hold.*

- (1) *If Assumption 1.2(a) holds, then  $y_t$  is geometrically ergodic and  $E_\pi(y_t^2) < \infty$ , where  $\pi$  is the stationary distribution of  $y_t$ .*  
 (2) *If  $E(e_t) = E(e_t^3) = 0$ ,  $E(e_t^4) < \infty$  and Assumption 1.2(b) holds, then  $y_t$  is geometrically ergodic and  $E_\pi(y_t^4) < \infty$ .*

**PROOF.** Since  $g_1$  and  $g_2$  are bounded on compacts,  $\phi$ -irreducibility and aperiodicity of  $Y_t$  is derived from Assumption 1.1. Moreover every compact set is a small set (see, e.g., Bhattacharya and Lee, 1995).

To prove part (1), we define a test function  $V : R^{p+q} \rightarrow R$  by

$$V(z_1, \dots, z_{p+q}) = \sum_{i=1}^{p+q} \gamma_i z_i^2 + 1, \quad (2.12)$$

where  $\gamma_i$ ,  $1 \leq i \leq p+q$ , are to be defined later. We first assume that  $p = q = 2$ . By (2.3)–(2.6) and simple calculation, we obtain that

$$\begin{aligned} |g_1(y_{t-1}, y_{t-2})| &\leq \theta_0 + \theta_1 |y_{t-1}| + \theta_2 |y_{t-2}|, \\ g_2^2(y_{t-1}, \dots, y_{t-4}) &\leq \sum_{i=1}^4 d_i y_{t-i}^2 + h_1(y_{t-1}, \dots, y_{t-4}), \end{aligned}$$

where

$$\begin{aligned} d_1 &= \alpha_1(1 + \theta_1 + \theta_2), \\ d_2 &= \alpha_1 \theta_1(1 + \theta_1 + \theta_2) + \alpha_2(1 + \theta_1 + \theta_2), \\ d_3 &= \alpha_1 \theta_2(1 + \theta_1 + \theta_2) + \alpha_2 \theta_1(1 + \theta_1 + \theta_2), \\ d_4 &= \alpha_2 \theta_2(1 + \theta_1 + \theta_2) \end{aligned}$$

and

$$\begin{aligned} h_1(y_{t-1}, \dots, y_{t-4}) &= \theta_0^2(\alpha_1 + \alpha_2) + 2\alpha_1 \theta_0(|y_{t-1}| + \theta_1 |y_{t-2}| + \theta_2 |y_{t-3}|) \\ &\quad + 2\alpha_2 \theta_0(|y_{t-2}| + \theta_1 |y_{t-3}| + \theta_2 |y_{t-4}|). \end{aligned}$$

For  $z = (z_1, \dots, z_4)$ , we have that

$$\begin{aligned}
 & E(V(Y_t)|Y_{t-1} = z) \\
 &= E\left[\gamma_1(g_1(z_1, z_2) + g_2(z_1, \dots, z_4)e_t)^2 + \sum_{i=2}^4 \gamma_i z_{i-1}^2 + 1 \mid Y_{t-1} = z\right] \\
 &\leq \gamma_1\left(\sum_{i=1}^2 \beta_i z_i^2 + \sigma^2 \sum_{i=1}^4 d_i z_i^2\right) + \sum_{i=2}^4 \gamma_i z_{i-1}^2 + 1 + h_1(z). \tag{2.13}
 \end{aligned}$$

Here  $\sum_{i=1}^2 \beta_i = (\sum_{i=1}^2 \theta_i)^2$ ,  $\sum_{i=1}^4 d_i = (1 + \sum_{i=1}^2 \theta_i)^2 (\sum_{i=1}^2 \alpha_i)$ . Use mathematical induction to obtain that, for arbitrary  $p, q \in \mathbb{Z}^+$ , and  $z = (z_1, \dots, z_{p+q})$ ,

$$g_1^2(z_1, \dots, z_p) \leq \sum_{i=1}^p \beta_i z_i^2 + o(\|z\|^2), \tag{2.14}$$

$$g_2^2(z_1, \dots, z_{p+q}) \leq \sum_{i=1}^{p+q} d_i z_i^2 + o(\|z\|^2) \tag{2.15}$$

with

$$\sum_{i=1}^p \beta_i = \left(\sum_{i=1}^p \theta_i\right)^2 \quad \text{and} \quad \sum_{i=1}^{p+q} d_i = \left(1 + \sum_{i=1}^p \theta_i\right)^2 \left(\sum_{i=1}^q \alpha_i\right). \tag{2.16}$$

Therefore

$$\begin{aligned}
 & E[V(Y_t)|Y_{t-1} = (z_1, \dots, z_{p+q})] \\
 &\leq \gamma_1\left(\sum_{i=1}^p \beta_i z_i^2 + \sigma^2 \sum_{i=1}^{p+q} d_i z_i^2\right) + \sum_{i=2}^{p+q} \gamma_i z_{i-1}^2 + o(\|z\|^2). \tag{2.17}
 \end{aligned}$$

Define  $\gamma_i$ ,  $i = 1, 2, \dots, p + q$ , by the same manner as given in (2.8) with  $n = p + q$ ,  $\sum \xi_i = \sum \beta_i + \sigma^2 \sum d_i$ . Then it follows from (2.17), Assumption 1.2(a) and Lemma 2.1 that there exists some constant  $r < 1$  such that

$$E[V(Y_t) \mid Y_{t-1} = z] \leq rV(z) + o(\|z\|^2),$$

from which inequalities (3.7) and (3.8) in Tjøstheim (1990) hold with some  $\epsilon > 0$ ,  $M < \infty$  and compact set  $K = \{\|z\| \leq k\}$  for sufficiently large  $k < \infty$ . Hence the geometric ergodicity and existence of the second moment of  $y_t$  follow (see, e.g., Meyn and Tweedie, 1993; An *et al.*, 1997).

To prove part (2), define  $V : R^{p+q} \rightarrow R$  by

$$V(z_1, \dots, z_{p+q}) = \sum_{i=1}^{p+q} \gamma_i z_i^4 + 1. \tag{2.18}$$

Note that

$$E(g_1 + g_2 e_t)^4 \leq (1 + 3\sigma^2)g_1^4 + (Ee_t^4 + 3\sigma^2)g_2^4, \quad (2.19)$$

$$g_1^4(z_1, \dots, z_p) = \left( \sum_{i=1}^p \theta_i z_i \right)^4 + h(z) \leq \sum_{i=1}^p \eta_i z_i^4 + h(z), \quad (2.20)$$

$$g_2^4 \leq \left( \sum_{i=1}^{p+q} d_i z_i^2 \right)^2 + h'(z) \leq \sum_{i=1}^{p+q} \delta_i z_i^4 + h'(z), \quad (2.21)$$

where  $\sum \eta_i = (\sum \theta_i)^4$ ,  $\sum \delta_i = (\sum d_i)^2 = (1 + \sum \theta_i)^4 (\sum \alpha_i)^2$ ,  $h(z)$  and  $h'(z)$  are functions satisfying  $h(z)/V(z)$ ,  $h'(z)/V(z) \rightarrow 0$  as  $\|z\| \rightarrow \infty$ .

It follows from (2.18)–(2.21) that

$$\begin{aligned} & E[V(Y_t) | Y_{t-1} = (z_1, \dots, z_{p+q})] \\ & \leq \gamma_1 [(1 + 3\sigma^2)g_1^4 + (Ee_t^4 + 3\sigma^2)g_2^4] + \sum_{i=2}^{p+q} \gamma_i z_{i-1}^4 + 1 \\ & \leq \gamma_1 \sum_{i=1}^{p+q} \left[ (1 + 3\sigma^2)\eta_i + (Ee_t^4 + 3\sigma^2)\delta_i \right] z_i^4 + \sum_{i=2}^{p+q} \gamma_i z_{i-1}^4 + f(z), \end{aligned} \quad (2.22)$$

where  $\eta_{p+1} = \dots = \eta_{p+q} = 0$  and  $f(z)/V(z) \rightarrow 0$  as  $\|z\| \rightarrow \infty$ . Since

$$\begin{aligned} & (1 + 3\sigma^2) \sum_{i=1}^{p+q} \eta_i + (Ee_t^4 + 3\sigma^2) \sum_{i=1}^{p+q} \delta_i \\ & = (1 + 3\sigma^2) \left( \sum_{i=1}^{p+q} \theta_i \right)^4 + (Ee_t^4 + 3\sigma^2) \left( 1 + \sum_{i=1}^{p+q} \theta_i \right)^4 \left( \sum_{i=1}^{p+q} \alpha_i \right)^2 \\ & < 1, \end{aligned}$$

applying Lemma 2.1 yields the conclusion.  $\square$

**EXAMPLE 1.** Suppose  $e_t \sim N(0, 1)$ . Then  $E(e_t) = E(e_t^3) = 0$ ,  $E(e_t^2) = 1$  and  $E(e_t^4) = 3$ . When  $\sum \theta_i = 0.5$ , then  $\sum \alpha_i \leq 0.149$  guarantees the geometric ergodicity and existence of 4<sup>th</sup> moment of  $y_t$ . If  $p = 0$ , i.e.,  $y_t$  is a threshold ARCH model, then  $\sum \alpha_i < 1$  ensures  $E(\varepsilon_t^2) < \infty$  and  $6(\sum \alpha_i)^2 < 1$ , i.e.,  $\sum \alpha_i \leq 0.4$  implies that  $E(\varepsilon_t^4) < \infty$ . Lemma 3.3 in An *et al.* (1997) emerges as a special case of Theorem 2.1(1).

**REMARK 1.** Recall that the geometric ergodicity of a Markov chain implies the absolute regularity and strongly mixing of the process.

3. STATIONARITY AND MOMENTS

In this section we consider the process  $y_t$  given by (1.1)–(1.3) *via* the different Markov chain. Let

$$W_t = (y_t, y_{t-1}, \dots, y_{t-p+1}, \varepsilon_t^2, \dots, \varepsilon_{t-q+1}^2). \tag{3.1}$$

It is an awkward problem to prove the irreducibility of  $W_t$ , which is a crucial step to prove the geometric ergodicity of  $W_t$ .

Throughout this section we assume that  $W_t$  is a Feller chain, that is, for each bounded continuous function  $f$  on  $R^{p+q}$ ,  $E[f(W_t)|W_{t-1} = w]$  is continuous in  $w$ . Note that if piecewise continuous functions  $\sum_{i=1}^r (\theta_{i0} + \sum_{j=1}^p \theta_{ij}y_{t-j})I_{it}$  and  $\sum_{k=1}^s (\alpha_{k0} + \sum_{j=1}^q \alpha_{kj}\varepsilon_{t-j}^2)J_{kt}$  are continuous, then  $W_t$  is a Feller chain.

The goal of this section is to find a region on which the process has a strictly stationary solution and has the finite  $m^{th}$  moment according to the stationary distribution. Many authors have studied the stationarity and moment conditions for nonlinear time series (*e.g.* Tweedie, 1988; Meyn and Tweedie, 1993; An *et al.*, 1997; Ling, 1999).

It is proved that if  $\max_i \sum_{j=1}^p |\theta_{ij}| < 1$  and  $\sum_{j=1}^q \max_i \{\alpha_{ij}\} < 1$ , then the strictly stationary solution  $y_t$  of (1.1)–(1.3) exists (see Lee and Kim, 2002). It is shown that  $E(e_t^{2m}) < \infty$ ,  $\sum_{j=1}^p \max_i |\theta_{ij}| < 1$  and  $\rho(\max_j E[\tilde{\alpha}_t^{(j)\otimes m}]) < 1$  ensure the existence of the  $m^{th}$  moment of  $y_t$ , where  $\rho(A)$  designates the spectral radius of a matrix  $A$  and the definition of  $\tilde{\alpha}_t^{(j)}$  is given in equation (4.6) in Ling (1999). But it is not simple to find a region of  $\alpha_{ij}$  on which  $\rho(\max_j E[\tilde{\alpha}_t^{(j)\otimes m}]) < 1$  is satisfied as well as the moment condition  $E(e_t^{2m}) < \infty$  on  $e_t$  can be weakened to obtain the finite  $m^{th}$  moment of  $y_t$ .

LEMMA 3.1. *If  $\sum_{i=1}^p \xi_i < 1$  and  $\sum_{i=1}^q \phi_i < 1$ ,  $n, m \in Z^+$ , then we can choose  $\gamma_i$ ,  $1 \leq i \leq p + q$  such that for some positive constant  $r < 1$ ,*

$$\begin{aligned} & \gamma_1 \left( \sum_{i=1}^p \xi_i z_i^m \right) + \sum_{i=2}^p \gamma_i z_{i-1}^m + \gamma_{p+1} \left( \sum_{i=1}^q \phi_i u_i^n \right) + \sum_{i=2}^q \gamma_{p+i} u_{i-1}^n \\ & \leq r \left( \sum_{i=1}^p \gamma_i z_i^m + \sum_{i=1}^q \gamma_{p+i} u_i^n \right). \end{aligned} \tag{3.2}$$

PROOF. Since  $\sum_{i=1}^p \xi_i < 1$  and  $\sum_{i=1}^q \phi_i < 1$ , choose  $\delta_1 > 0$  and  $\delta_2 > 0$  by  $\sum \xi_i + \delta_1 = 1$  and  $\sum \phi_i + \delta_2 = 1$ . Now choose  $\gamma_1 > 0$  and  $\gamma_{p+1} > 0$  arbitrary and

define  $\gamma_i$ ,  $2 \leq i \leq p$ ,  $p+2 \leq i \leq p+q$ , by

$$\gamma_{i+1} = \gamma_1 \left( 1 - \zeta_1 - \cdots - \xi_i - \frac{i\delta}{p} \right), \quad 1 \leq i \leq p-1, \quad (3.3)$$

$$\gamma_{p+1+i} = \gamma_{p+1} \left( 1 - \phi_1 - \cdots - \phi_i - \frac{i\delta}{q} \right), \quad 1 \leq i \leq q-1. \quad (3.4)$$

Then by the same reason adopted in the proof of Lemma 2.1, (3.2) holds with  $r = \max\{1 - \delta_1/p, 1 - \delta_2/q\}$ .  $\square$

**THEOREM 3.1.** *Let  $m \in Z^+$  and  $n = [m/2] + 1$  where  $[\cdot]$  denotes the Gaussian number. Suppose that  $E(e_t^{2n}) < \infty$ . If  $\sum \theta_i < 1$  and  $E(e_t^{2n})(\sum \alpha_i)^n < 1$ , then  $E_\pi(|y_t|^m) < \infty$ .*

**PROOF.** To avoid the computational complexity, we assume that  $p = q = 2$ . The case  $p > 2$  or  $q > 2$  is entirely analogous, but involves messier notation. Define a test function

$$V(z_1, \dots, z_p, u_1, \dots, u_q) = \sum_{i=1}^p \gamma_i |z_i|^m + \sum_{i=1}^q \gamma_{p+i} u_i^n, \quad (3.5)$$

where  $\gamma_i$  is to be defined by the equations (3.3) and (3.4).

For  $(z, u) = (z_1, z_2, u_1, u_2)$ ,

$$\begin{aligned} & E[V(Y_t) | Y_{t-1} = (z, u)] \\ &= E[\gamma_1(\theta_0 + \theta_1|z_1| + \theta_2|z_2| + (\alpha_0 + \alpha_1 u_1 + \alpha_2 u_2)^{1/2} e_t)^m] + \gamma_2 |z_1|^m \\ &\quad + \gamma_3 E(e_t^{2n})(\alpha_0 + \alpha_1 u_1 + \alpha_2 u_2)^n + \gamma_4 u_1^n \\ &= \gamma_1(\theta_1|z_1| + \theta_2|z_2|)^m + \gamma_2 |z_1|^m + \gamma_3 E(e_t^{2n})(\alpha_1 u_1 + \alpha_2 u_2)^n + \gamma_4 u_1^n + f(z, u) \\ &\leq \gamma_1(\beta_1|z_1|^m + \beta_2|z_2|^m) + \gamma_2 |z_1|^m + \gamma_3 E(e_t^{2n})(\eta_1 u_1^n + \eta_2 u_2^n) + \gamma_4 u_1^n + f(z, u). \end{aligned}$$

Here

$$\frac{f(z, u)}{V(z, u)} \rightarrow 0 \quad \text{as } V(z, u) \rightarrow \infty,$$

for  $|z_i|^k u_j^{(m-k)/2} \leq (|z_i|^m + u_j^n)^{k/m + (m-k)/(2n)}$  and  $k/m + (m-k)/(2n) < 1$ ,  $1 \leq k \leq m-1$ .

Since  $\beta_1 + \beta_2 = (\theta_1 + \theta_2)^m$  and  $\eta_1 + \eta_2 = (\alpha_1 + \alpha_2)^n$ , from assumptions  $\sum \theta_i < 1$  and  $E(e_t^{2n})(\sum \alpha_i)^n < 1$  and Lemma 3.1, we obtain the existence of  $r < r' < 1$  such that

$$E[V(Y_t) | Y_{t-1} = (z, u)] \leq rV(z, u) + f(z, u) \leq r'V(z, u) - \epsilon, \quad \|z\| > M \quad (3.6)$$



for some  $\epsilon > 0$  and  $M < \infty$ . Conclusion can be deduced from Theorem 2 of Tweedie (1988).  $\square$

EXAMPLE 2. Suppose that  $e_t \sim N(0, 1)$ . If  $m = 4$ , then  $n = 3$  and  $E(e_t^6) = 15$ . Therefore  $E|y_t|^4 < \infty$  provided that  $\sum \theta_i < 1$  and  $\sum \alpha_i \leq 0.4$ .

REMARK 2. Let  $y_t$  be generated by (1.1)–(1.3) with  $r = s = 2$ ,  $I_{1t} = I(y_t - y_{t-b} \geq \lambda)$ ,  $I_{2t} = 1 - I_{1t}$  and  $J_{1t} = I(y_t - y_{t-d} \geq \lambda')$ ,  $J_{2t} = 1 - J_{1t}$  for some constants  $\lambda, \lambda' \in R^1$ . Such a model is called a momentum threshold AR-ARCH model (MTARCH). We can easily show that Theorem 3.1 still holds for the MTARCH process  $\{y_t\}$ . The proof for MTARCH process follows essentially the same line of Theorem 3.1.

#### REFERENCES

- AN, H., CHEN, M. AND HUANG, F. (1997). "The geometric ergodicity and existence of moments for a class of nonlinear time series model", *Statistics & Probability Letters*, **31**, 213–224.
- BHATTACHARYA, R. N. AND LEE, C. (1995). "On geometric ergodicity of nonlinear autoregressive models", *Statistics & Probability Letters*, **22**, 311–315.
- BOLLERSLEV, T., CHOU, R. T. AND KRONER, K. F. (1992). "ARCH modeling in finance", *Journal of Econometrics*, **52**, 5–59.
- CHAN, K. S. (1993). "Consistency and limiting distribution of least squares estimation of a threshold autoregressive model", *The Annals of Statistics*, **21**, 520–533.
- ENGLE, R. F. (1982). "Autoregressive conditional heteroscedasticity with estimates of the variance of the United Kingdom inflation", *Econometrica*, **50**, 987–1007.
- ENGLE, R. F. AND BOLLERSLEV, T. (1986). "Modeling the persistence of conditional variances", *Econometric Reviews*, **5**, 1–87.
- GUÉGAN, D. AND DIEBOLT, J. (1994). "Probabilistic properties of the  $\beta$ -ARCH model", *Statistica Sinica*, **4**, 71–87.
- HANSEN, B. E. (1999). "Threshold effects in non-dynamic panels : Estimation, testing, and inference", *Journal of Econometrics*, **93**, 345–368.
- HANSEN, B. E. (2000). "Sampling splitting and threshold estimation", *Econometrica*, **68**, 575–603.
- HWANG, S. Y. AND WOO, M. (2001). "Threshold ARCH(1) processes : Asymptotic inference", *Statistics & Probability Letters*, **53**, 11–20.
- LEE, O. (2002). "On strict stationarity of nonlinear ARMA processes with nonlinear GARCH innovations", preprint.
- LEE, O. AND KIM, M. (2002). "On stationarity of nonlinear AR processes with nonlinear ARCH errors", *Communications of the Korean Mathematical Society*, **17**, 309–319.
- LI, C. W. AND LI, W. K. (1996). "On a double threshold autoregressive conditional heteroscedastic time series model", *Journal of Applied Econometrics*, **11**, 253–274.
- LING, S. (1999). "On the probabilistic properties of a double threshold ARMA conditional heteroskedastic model", *Journal of Applied Probability*, **36**, 688–705.

- LU, Z. (1996). "A note on geometric ergodicity of autoregressive conditional heteroscedasticity (ARCH) model", *Statistics & Probability Letters*, **30**, 305–311.
- MASRY, E. AND TJØSTHEIM, D. (1995). "Nonparametric estimation and identification of nonlinear ARCH time series", *Econometric Theory*, **11**, 258–289.
- MEYN, S. P. AND TWEEDIE, R. L. (1993). *Markov Chains and Stochastic Stability*, Springer, London.
- NELSON, D. B. (1991). "Conditional heteroscedasticity in asset returns : a new approach", *Econometrica*, **59**, 347–370.
- SCHWERT, G. W. (1989). "Why do stock market volatility change over time?", *The Journal of Finance*, **44**, 1115–1153.
- TJØSTHEIM, D. (1990). "Nonlinear time series and Markov chains", *Advances in Applied Probability*, **22**, 587–611.
- TONG, H. (1978). "On a threshold model", In *Pattern Recognition and Signal Processing* (C. H. Chen, ed.), Sijthoff and Noordhoff, Amsterdam.
- TONG, H. (1990). *Nonlinear Time Series : A Dynamical System Approach*, Oxford University Press, Oxford.
- TONG, H. AND LIM, K. S. (1980). "Threshold autoregressive, limit cycles and cyclical data", *Journal of the Royal Statistical Society*, **B42**, 245–292.
- TWEEDIE, R. L. (1988). "Invariant Measures for Markov chains with no irreducibility assumptions", *Journal of Applied Probability*, **25A**, 275–285.
- WONG, H. AND LI, W. K. (1997). "On a multivariate conditional heteroscedastic model", *Biometrika*, **84**, 111–123.